# Autopolyploid establishment through polygenic adaptation

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#### Abstract

We define the infinitesimal model of quantitative genetics for the inheritance of 6 an additive quantitative trait in a mixed-ploidy population consisting of diploid, 7 triploid and autotetraploid individuals producing haploid and diploid gametes. We 8 implement efficient simulation methods and use these to study the quantitative 9 genetics of mixed-ploidy populations and the establishment of autotetraploids in a 10 new habitat. We show that, when migration from the source population is rare, 11 autotetraploids are more likely to establish in the new habitat than diploids under 12 a very broad range of conditions, but that this is unlikely to sufficiently counter the 13 scarcity of tetraploid founders when the source is predominantly diploid. We assess 14 in more detail how minority cytotype exclusion interacts with migration load in the 15 establishment process and evaluate the impact of additional sources of prezygotic 16 isolation, specifically selfing and assortative mating, on the relative establishment 17 probabilities of the different cytotypes in the presence of maladaptive migration. 18 In the discussion we consider how inbreeding depression may impact our findings. 19

Keywords: polyploidy, adaptation, establishment, inbreeding, quantitative genetics

## <sup>22</sup> Introduction

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Many plant species exhibit ploidy variation (Levin, 2002; Soltis et al., 2007; Rice et al., 2015), and many of these *mixed-ploidy* species have populations in which different cytotypes coexist or form contact zones (Kolář et al., 2017). How such mixed-ploidy populations emerge and are maintained has proven somewhat challenging to understand.

Consider for instance a randomly mating diploid population. Under the commonly 27 accepted view that new polyploid plants are mostly formed through the union of unre-28 duced gametes (Bretagnolle and Thompson, 1995; Herben et al., 2016; Kreiner et al., 29 2017b), a new tetraploid individual originating by a chance encounter of two unreduced 30 diploid gametes (an event occurring at an appreciable rate; Kreiner et al. (2017a)) is 31 highly unlikely to establish a stable tetraploid subpopulation, as most of its gametes will 32 end up in unfit hybrids of odd ploidy level (a phenomenon referred to as 'triploid block', 33 see Ramsey and Schemske (1998); Köhler et al. (2010); Brown et al. (2024)). This posi-34 tive frequency dependence effect in mixed-ploidy populations is commonly referred to as 35

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minority cytotype exclusion (MCE), after Levin (1975). It is well-appreciated that, as a
consequence of MCE, the rate of unreduced gamete formation needs to be extraordinarily
high for tetraploids to establish a stable subpopulation in a large random mating population initially dominated by diploids (Felber and Bever (1997), see also supplementary
section S2.1).

Hence, to explain the widespread occurrence of mixed-ploidy populations, additional 41 factors besides the continuous formation of polyploids through the union of unreduced 42 gametes need to be considered. Firstly, chance establishment of tetraploids through drift 43 could occur. Indeed, the problem is somewhat analogous to the spread of underdomi-44 nant chromosomal rearrangements, where local establishment through genetic drift and 45 subsequent spreading in a subdivided population by means of local extinction and re-46 colonization has been suggested as a plausible model (Lande, 1985). However, MCE is 47 quite strong under random mating, and the population size has to be very small for local 48 tetraploid establishment to occur at an appreciable rate (Rausch and Morgan (2005), 49 see also supplementary section S2.2). Secondly, any form of *prezyqotic isolation* between 50 cytotypes could promote establishment of polyploid cytotypes by alleviating MCE. Par-51 ticularly relevant are assortative mating by cytotype (for instance through phenological 52 differences across cytotypes, or differences in pollinators; Husband and Sabara (2004); 53 Kolář et al. (2017)), self-fertilization (Rausch and Morgan, 2005; Novikova et al., 2023), lo-54 calized dispersal (Baack, 2005; Kolář et al., 2017) and asexual reproduction (Van Drunen 55 and Friedman, 2022). Finally, selection may be invoked to explain the establishment of 56 polyploids. Tetraploids may have higher relative fitness than their diploid counterparts 57 due to reduced inbreeding depression (Husband and Schemske, 1997; Ronfort, 1999; Otto 58 and Whitton, 2000; Husband and Sabara, 2004; Husband et al., 2008; Clo and Kolář, 59 2022), or due to being better adapted to (changing) environmental conditions (Van de 60 Peer et al., 2021). However, none of these factors is likely to explain by itself the estab-61 lishment of polyploids, and the consensus in the field appears to be that some mix of the 62 above is required to explain the occurrence of mixed-ploidy populations in nature (Kolář 63 et al., 2017; Mortier et al., 2024). 64

Many empirical studies of mixed-ploidy populations find that polyploids established 65 in peripheral habitats at the edge of a species' range (reviewed in Griswold (2021)), 66 and this is in accord with large scale biogeographical patterns (Rice et al., 2019). This 67 raises the question: which aspects of the process of adaptation to marginal habitats 68 could promote the establishment of polyploid populations? In a peripheral habitat, a 69 new polyploid population may be more likely to reach an appreciable size as it evades to 70 some extent the negative effects of MCE (Levin, 1975). However, at the same time, such 71 peripheral habitats are likely to present adaptive challenges to establishment (Kawecki, 72 2008; Sachdeva et al., 2022), and if polyploids are able to colonize such habitats at an 73 appreciable rate, they must somehow be better adapted to local conditions, or more able 74 to adapt to those conditions despite inbreeding and maladaptive migration, compared to 75 diploids. 76

More often than not, local adaptation is polygenic in nature (Pritchard and Di Rienzo,
2010; Barghi et al., 2020; Bomblies and Peichel, 2022), involving many weakly selected
variants across the genome, and adaptation during polyploid establishment in a marginal
habitat is unlikely to present an exception. Recent studies on local adaptation in autopolyploids indeed tend to find a polygenic basis of adaptation (Bohutínská et al., 2021;
Konečná et al., 2021, 2022), however it is not clear how observed adaptive differentiation
in established tetraploid populations relates to adaptation that may have occurred during

<sup>84</sup> initial establishment.

While there have been substantial modeling efforts aimed at understanding autotetraploid establishment within diploid populations (Levin, 1975; Felber, 1991; Felber and Bever, 1997; Rausch and Morgan, 2005; Oswald and Nuismer, 2011; Clo et al., 2022), the problem of polyploid establishment in peripheral habitats remains largely unaddressed (but see Griswold (2021)), despite its centrality to verbal arguments about the establishment of polyploids in natural populations (Kolář et al., 2017; Van de Peer et al., 2021; Clo, 2022b). Here we develop a model for the establishment of a mixed-ploidy population in a

92 novel, unoccupied habitat based on Barton and Etheridge (2018). In order to establish 93 in the novel habitat, the population has to adapt to local environmental conditions. We 94 assume fitness is determined by directional selection on a single polygenic trait, which 95 can be interpreted as log fitness at low density in the new habitat. As in Barton and 96 Etheridge (2018), we assume the trait follows the infinitesimal model (sensu Barton et al. 97 (2017), i.e. the 'Gaussian descendants' infinitesimal model (Turelli, 2017)). We extend 98 the infinitesimal model, and the approach for exact simulation of trait evolution under 99 the infinitesimal model, to mixed-ploidy populations. This is a first contribution of the 100 present paper. We then use simulations to study tetraploid establishment, both from 101 single migrants and under continuous migration from a predominantly diploid source 102 population, examining the effects of autopolyploid genetics, maladaptive migration, self-103 ing and assortative mating on the probability that autotetraploids establish in the novel 104 habitat. 105

### <sup>106</sup> Model and Methods

#### <sup>107</sup> Mixed-ploidy population model

Our notation is summarized in table 1. We consider a mixed-ploidy population of size Nconsisting of  $N_2$  diploid,  $N_3$  triploid and  $N_4 = N - N_2 - N_3$  tetraploid individuals. We assume an individual of ploidy level k forms haploid and diploid gametes with proportions  $u_{k1}$  and  $u_{k2}$ , as well as a proportion  $1 - u_{k1} - u_{k2}$  inviable (e.g. aneuploid or polyploid) gametes. The (relative) fecundity of a k-ploid individual is hence  $u_{k1} + u_{k2}$ . Unless stated otherwise, we will assume

$$\begin{pmatrix} u_{21} & u_{22} \\ u_{31} & u_{32} \\ u_{41} & u_{42} \end{pmatrix} = \begin{pmatrix} 1-u & u \\ v & v \\ 0 & 1-u \end{pmatrix}$$
(1)

where u is referred to as the proportion of unreduced gametes, and 2v is the proportion of euploid gametes produced by a triploid individual. The rate of unreduced gamete production is hence fixed across individuals, and is assumed to be the same in diploids and tetraploids.

<sup>118</sup> When two individuals mate, we assume they produce gametes according to their ploidy <sup>119</sup> level (eq. (1)), which randomly combine to produce offspring (which may be inviable if one <sup>120</sup> of the contributing gametes is inviable). Intrinsic fitness disadvantages associated with <sup>121</sup> particular zygotic ploidy levels or cross types (e.g. modeling phenomena such as 'triploid <sup>122</sup> block') can be straighforwardly included at this level. An analysis of a deterministic model <sup>123</sup> (i.e. where  $N \to \infty$ ) for the cytotype dynamics and equilibrium cytotype composition <sup>124</sup> under random mating is included in supplementary section S2.1 (see also Felber and

notation	description
N	total population size
$N_k$	population size of the $k$ -ploid cytotype
$\pi_k$	deterministic equilibrium frequency of the $k$ -ploid cytotype
u	probability of unreduced gamete formation $(u = u_{22} = 1 - u_{42})$
v	probability that a triploid produces a haploid/diploid gamete ( $v = u_{31} = u_{32}$ )
m	expected number of migrants per generation arriving in the new habitat
$z_i$	trait value of individual $i$
$c_i$	ploidy level of individual $i$
$g_i$	ploidy level of gamete produced by individual $i$ in a particular cross
V	segregation variance in the reference diploid population
$V_{i,k}$	gametic segregation variance associated with the production of a $k$ -ploid gamete
	by individual $i$
$\mathcal{V}_k$	genetic variance associated with a haploid genome in the $k$ -ploid reference pop-
	ulation (i.e. a $k$ -ploid non-inbred population at HWLE)
$\beta_k$	scaling factor for allelic effects in $k$ -ploids
$F_i$	inbreeding coefficient in individual $i$
$\Phi_{ij}$	coancestry coefficient for individuals $i$ and $j$
$\alpha_k$	probability that the two genes at a locus in a diploid gamete formed by a $k$ -ploid
	individual descend from the same parental gene copy
$\gamma$	strength of directional selection in the new habitat
heta	trait value beyond which the growth rate becomes positive in the new habitat
$w_{ij}$	fitness of parental pair $(i, j)$ (expected zygotic fitness of offspring of pair $(i, j)$ )
$w_{ij}^{kl}$	expected fitness of zygote from parental pair $(i, j)$ when i contributes a k-ploid
	gamete and $j$ contributes a $l$ -ploid gamete
$w_i$	expected number of offspring from individual $i$ surviving to reproductive age
$\sigma_k$	rate of self-fertilization in $k$ -ploids
$ ho_k$	probability of assortative mating in $k$ -ploids

Table 1: Glossary of the notation used in the main text.

Bever (1997); Kauai et al. (2024)). The stochastic version for finite and constant N is analyzed briefly in supplementary section S2.2.

#### 127 Infinitesimal model

The basic infinitesimal model. Consider a population which expresses a quantitative trait determined by a large number of additive loci of small effect. The infinitesimal model approximates the inheritance of such a trait by assuming that the trait value  $Z_{ij}$  of a random offspring from parents with trait values  $z_i$  and  $z_j$  follows a Gaussian distribution with mean equal to the midparent value and variance which is independent of the mean:

$$Z_{ij} \sim \mathcal{N}\left(\frac{z_i + z_j}{2}, V_{ij}\right)$$
 (2)

Here,  $V_{ij}$  is referred to as the *segregation variance* in family (i, j). This is the variation generated among offspring from the same parental pair due to random Mendelian segregation in meiosis. This approximation can be justified as arising from the limit where the number of loci determining the trait tends to infinity (Barton et al., 2017).

An alternative, and for our purposes useful, way to characterize the model is to write  $Z_{ij} = Y_i + Y_j$ , where  $Y_i$  and  $Y_j$  are independent Gaussian random variables  $Y_i \sim \mathcal{N}\left(\frac{z_i}{2}, V_i\right)$  (and similarly for  $Y_j$ ). We refer to  $Y_i$  as the (random) gametic value of individual *i*, and to  $V_i$  as the gametic segregation variance of individual *i*. This formulation is helpful in that it highlights that Mendelian segregation occurs independently in both parents to produce gametes, which then combine additively to determine the offspring trait value. This model applies readily to an autopolyploid population expressing a trait with infinitesimal genetics. However, the segregation variance will be determined by the details of tetraploid meiosis, which differ from those of diploid meiosis (see below).

In a finite population, the segregation variance will decay over time as the population becomes more inbred (Mendelian segregation at homozygous loci does not generate any variation). When  $F_i$  is the inbreeding coefficient relative to some ancestral reference population with gametic segregation variance V (i.e. the probability that two genes at a locus in individual i sampled without replacement are identical by descent), the gametic segregation variance of individual i will be  $V_i = (1 - F_i)V$ . This holds for both diploids and tetraploids (supplementary section S2.5.1, also Moody et al. (1993)).

Scaling of traits across ploidy levels. If we would naively assume that the allelic 148 effects underlying an additive trait are identical across ploidy levels, a tetraploid offspring 149 from a cross between two diploids would have, on average, a trait value which is the sum 150 of the parental trait values. This is not likely to reflect biological reality: tetraploids 151 do not tend to have, for instance, twice the size of their diploid progenitors on average 152 (e.g. Porturas et al. (2019)). Furthermore, the genetic variance at Hardy-Weinberg and 153 linkage equilibrium (HWLE) in a large non-inbred tetraploid population will be twice 154 that of their diploid counterparts under such assumptions, which is similarly unrealistic 155 (Clo, 2022a). 156

In order to account for this, we introduce a scaling factor  $\beta_k$ , accounting for the effects of polyploidization *per se* on trait expression in *k*-ploids. To introduce and interpret this parameter, we consider an *L*-locus additive model, with two alleles (0 and 1) at each locus. For a *k*-ploid individual, let  $X_{i,j}$  be the allele at homolog *j* of locus *i*. We assume the trait value is determined by

$$z = \sum_{i=1}^{L} \sum_{j=1}^{k} a_{i,k} X_{i,j}$$
(3)

Where  $a_{i,k}$  is the allelic effect of the 1 allele at locus *i* in *k*-ploids. The genetic variance at HWLE in *k*-ploids  $(\tilde{V}_{z,k})$  will then be

$$\tilde{V}_{z,k} = k \sum_{i=1}^{L} a_{i,k}^2 p_i q_i = k \mathcal{V}_k \tag{4}$$

where we refer to  $\mathcal{V}_k$  as the variance associated with a haploid genome in k-ploids at HWLE. If we now assume  $a_{i,k} = \beta_k a_{i,2}$ , i.e. allelic effects in k-ploids are as in diploids, but scaled homogeneously by a factor  $\beta_k$ , and assume equal allele frequencies in the different cytotypes, we will have

$$\frac{\tilde{V}_{z,k}}{\tilde{V}_{z,2}} = \frac{k\mathcal{V}_k}{2\mathcal{V}_2} = \frac{k}{2}\beta_k^2 \tag{5}$$

Note that by definition  $\beta_2 = 1$ . Under the infinitesimal model (where  $a_{i,2} \to 0$  as  $L \to \infty$ ), we have  $\tilde{V}_{z,2} = 2\mathcal{V}_2 = 2V$  (Barton et al., 2017), where V is the segregation variance in the diploid population. Hence, in the infinitesimal limit we have  $\mathcal{V}_k = \beta_k^2 \mathcal{V}_2 = \beta_k^2 V$ .

**Table 2:** Gametic segregation variance for haploid and diploid gametes produced by the three cytotypes in the mixed-ploidy model.  $F_i$  is the inbreeding coefficient of individual *i* (producing the gamete), whereas  $\alpha_k$  is the probability that a diploid gamete from a *k*-ploid individual receives two copies of the same parental gene. We assume  $\alpha_3 \leq 1/4$  and  $\alpha_4 \leq 1/6$  (see supplementary section S2.6).

cytotype	haploid gamete variance	diploid gamete variance
diploid	$\frac{1}{2}(1-F_i)V$	$2\alpha_2(1-F_i)V$
triploid	$\frac{\overline{2}}{\overline{3}}(1-F_i)V$	$\frac{2}{3}(1+3\alpha_3)(1-F_i)V$
tetraploid	_	$(1+2\alpha_4)(1-F_i)V$

Mixed-ploidy infinitesimal model. We can extend the infinitesimal model to the mixed-ploidy case, assuming that the gametic value, on the diploid trait scale, associated with a k-ploid gamete ( $k \in \{1, 2\}$ ) from individual *i* of ploidy level  $c_i \in \{2, 3, 4\}$  is a Gaussian random variable  $Y_{i,k}$  with distribution

$$Y_{i,k} \sim \mathcal{N}\left(\frac{k}{c_i} \frac{z_i}{\beta_{c_i}}, V_{i,k}\right) \tag{6}$$

where  $V_{i,k}$  is the gametic segregation variance associated with the production of a k-ploid gamete by individual *i* (see below). The trait value of an individual originating from the union of a k-ploid gamete of individual *i* and an *l*-ploid gamete from individual *j* is then

$$Z_{ij}^{kl} = \beta_{k+l} \left( Y_{i,k} + Y_{j,l} \right)$$

i.e.,  $Z_{ij}^{kl}$  is a Gaussian random variate with distribution  $Z_{ij}^{kl} \sim \mathcal{N}\left(\overline{z_{ij}^{kl}}, V_{ij}^{kl}\right)$  where

$$\overline{z_{ij}^{kl}} = \beta_{k+l} \left( \frac{k}{c_i} \frac{z_i}{\beta_{c_i}} + \frac{l}{c_j} \frac{z_j}{\beta_{c_j}} \right)$$
$$V_{ij}^{kl} = \beta_{k+l}^2 (V_{i,k} + V_{j,l})$$
(7)

The gametic segregation variance associated with the production of diploid gametes 175 depends not only on the segregation variance in the base population (V) and the in-176 breeding coefficient (F), but also on the detailed assumptions of how the meiotic process 177 takes place. Importantly however, the latter only affect the gametic segregation variance 178 through the quantity  $\alpha_k$ , which is the probability that a k-ploid transmits two copies of the 179 same homolog to a diploid gamete. Note that  $\alpha_4$ , the probability that a diploid gamete of 180 a tetraploid individual carries two copies of the same homolog, is the probability of *double* 181 reduction (e.g. Lynch and Walsh (1998) p.57), and is upper bounded by 1/6 (Stift et al., 182 2008). The value of  $\alpha_2$  depends on the relative frequency of unreduced gamete formation 183 through so-called *first* and *second division restitution* (Bretagnolle and Thompson, 1995; 184 De Storme and Geelen, 2013). We summarize the expressions for the gametic segregation 185 variance in table 2. Detailed derivations can be found in supplementary section S2.6. 186

Writing  $\bar{z}_2$  for the mean trait value in the diploid subpopulation, eq. (7) implies that 187 a tetraploid offspring from a random diploid parental pair has an expected trait value 188 equal to  $\bar{z}_4 = 2\beta_4 \bar{z}_2$ . This hence implies that  $|\bar{z}_4| \ge |\bar{z}_2|$ , with equality only when  $\beta_4 = 1/2$ 189 or  $\bar{z}_2 = \bar{z}_4 = 0$ . In other words, when  $\beta_4 > 1/2$ , we would have for all but  $\bar{z}_2 = 0$  that 190 newly formed tetraploids have more extreme phenotypes on average than their diploid 191 parents. In our establishment model (see below), we shall therefore always consider the 192 case where  $\bar{z}_2 = 0$  in the source population, and think of the trait value modeled as the 193 deviation from the mean phenotype in the source population. 194

The property that polyploid phenotypes are more extreme on average makes sense if 195 we consider the underlying genetic model: if we assume the source population is at an 196 equilibrium between mutation, drift and stabilizing selection, segregating genetic variants 197 that affect the trait will be equally likely to have positive or negative allelic effects (Hay-198 ward and Sella, 2022). At equilibrium, any diploid individual with trait value  $z - \bar{z}_2 > 0$ 199 then carries an excess of variants with positive allelic effects on the trait, and this excess 200 should be exaggerated in its polyploid offspring, which carry *twice* the excess of their 201 parent on average. In other words, the mixed-ploidy model is not 'coordinate-free' as 202 the basic infinitesimal model is: eq. (7) only makes sense when the trait values that are 203 modeled correspond to deviations from the mean values associated with an underlying 204 equilibrium state. 205

**Recursions for inbreeding coefficients** We can simulate the mixed-ploidy infinitesimal model for a finite population through a straightforward extension of the approach outlined in Barton et al. (2017), provided we can efficiently track inbreeding and coancestry coefficients across the different ploidy levels. Denoting the parents of individual i by k and l, the recursion for the inbreeding coefficients in the mixed-ploidy case becomes

$$F_{i} = \Phi_{kl} \qquad \text{if } c_{i} = 2$$

$$F_{i} = \frac{1}{3} \left( F_{k}^{*} + 2\Phi_{kl} \right) \qquad \text{if } c_{i} = 3, g_{k} = 2, g_{l} = 1$$

$$F_{i} = \frac{1}{3} \left( F_{l}^{*} + 2\Phi_{kl} \right) \qquad \text{if } c_{i} = 3, g_{k} = 1, g_{l} = 2$$

$$F_{i} = \frac{1}{6} \left( F_{k}^{*} + F_{l}^{*} + 4\Phi_{kl} \right) \qquad \text{if } c_{i} = 4 \qquad (8)$$

where  $F_k^* = \alpha_{c_k} + (1 - \alpha_{c_k})F_k$  (supplementary section S2.5.1). The recursion for the coancestry coefficients is given by

$$\Phi_{ii} = \frac{1}{c_i} \left( 1 + (c_i - 1)F_i \right)$$
  
$$\Phi_{ij} = \sum_k \sum_l P_{ik} P_{jl} \Phi_{kl} \qquad i \neq j \qquad (9)$$

where the sums are over individuals in the parental population, and where  $P_{ik} \in \{0, \frac{1}{3}, \frac{1}{2}, \frac{2}{3}, 1\}$ is the probability that a gene copy in individual *i* is derived from parent *k*.

#### 208 Establishment model

Our model for the establishment of a population in an initially unoccupied habitat is based on Barton and Etheridge (2018). We assume a large non-inbred 'mainland' mixedploidy population at HWLE and cytotype equilibrium, with  $\mathbb{E}[z] = 0$  irrespective of the cytotype. The equilibrium trait value distribution for the different cytotypes on the mainland is complicated in general, but a very accurate approximation (which we use throughout) is readily obtained for the case where u is small (see supplementary section S2.7).

In generation t,  $M(t) \sim \text{Poisson}(m)$  migrant individuals arrive on an island (the new habitat) joining  $N^*(t)$  resident individuals. We assume  $N^*(0) = 0$  unless stated otherwise. The migrant individuals are assumed to be unrelated to the resident individuals.

After migration there are  $N(t) = N^*(t) + M(t)$  individuals on the island. These indi-219 viduals reproduce sexually by random mating (allowing random selfing, see below for 220 modifications involving partial selfing and assortative mating) to produce a large (effec-221 tively infinite) number of zygotes, of which  $N^*(t+1)$  survive to reproductive age with 222 a probability proportional to individual fitness, where  $N^*(t+1)$  is Poisson distributed 223 with mean  $N(t)\overline{W}$  and  $\overline{W}$  is the mean fitness among zygotes. We assume the trait to 224 be under directional selection, with the fitness of a zygote with trait value z given by 225  $w(z) = e^{\gamma(z-\theta)}$ , where  $\gamma$  is the intensity of directional selection and  $\theta$  is the trait value 226 for which the growth rate of the island population becomes positive. As the popula-227 tion becomes better adapted, eventually, some form of density regulation must limit its 228 growth. We ignore density regulation throughout and focus on the initial establishment 229 phase, defining establishment as reaching a population of size 100. Note that this model 230 corresponds to a life cycle where selection occurs before migration within a generation. 231

Again following Barton and Etheridge (2018), we simulate the model by first calculating the fitness of each parental pair (i, j), which is defined as the expected fitness of a random zygote derived from this pair

$$w_{ij} = \sum_{k=1}^{2} \sum_{l=1}^{2} w_{ij}^{kl} = \sum_{k=1}^{2} \sum_{l=1}^{2} u_{c_i,k} u_{c_j,l} \mathbb{E}\left[e^{\gamma(Z_{ij}^{kl} - \theta)}\right]$$
(10)

The expected number of offspring individuals of an individual *i* surviving to reproductive age is then  $w_i = \frac{1}{N(t)} \sum_j w_{ij}$ . The expectation on the right hand side of eq. (10) can be calculated from eq. (7) using the moment-generating function of the Gaussian, i.e.

$$\mathbb{E}\left[e^{\gamma(Z_{ij}^{kl}-\theta)}\right] = e^{\gamma\left(\overline{z_{ij}^{kl}}-\theta\right) + \frac{\gamma^2}{2}V_{ij}^{kl}}$$
(11)

Having calculated the  $w_{ij}$ , the number of offspring surviving into the next generation  $N^*(t+1)$  is sampled from a Poisson distribution with mean  $N(t)\overline{W} = \sum_i w_i = \sum_{i,j} w_{ij}/N(t)$ . Offspring cytotypes and trait values are then determined by sampling parental pairs and gametes proportional to  $w_{ij}^{kl}$ , and sampling a trait value from the offspring trait value distribution conditional on survival to reproductive age, which is Gaussian with mean  $\overline{z_{ij}^{kl}} + \gamma V_{ij}^{kl}$  and variance  $V_{ij}^{kl}$ .

#### <sup>244</sup> Self-fertilization and assortative mating

We model partial self-fertilization by assuming that a proportion  $\sigma_{c_i}$  of the ovules of individual *i* with ploidy level  $c_i$  are fertilized by self-pollen, while the remaining proportion  $1 - \sigma_{c_i}$  are fertilized by randomly sampled pollen (which may be self-pollen with probability 1/N). That is, the expected number of offspring from individual *i* as mother surviving after selection is

$$w_{i} = \sigma_{c_{i}} w_{ii} + (1 - \sigma_{c_{i}}) \left[ \frac{1}{N} \sum_{j=1}^{N} w_{ij} \right]$$
(12)

<sup>250</sup> We hence assume no pollen limitation (all outcrossing ovules are fertilized), and no pollen

discounting (the probability of being a father is unaffected by an individual's selfing rate).

<sup>252</sup> When modeling self-incompatibility, we assume there is no intrinsic disadvantage to self-

<sup>253</sup> incompatibility, except when there is only a single individual in the population, i.e.

$$w_{i} = \begin{cases} \frac{1}{N-1} \sum_{j \neq i} w_{ij} & \text{if } N > 1\\ 0 & \text{if } N = 1 \end{cases}$$
(13)



Figure 1: The infinitesimal model in autotetraploids. Comparisons are shown for the decay of the genetic variance  $(V_z)$  due to inbreeding in exact simulations of the infinitesimal model in autotetraploids against individual-based simulations of autotetraploid populations with L. unlinked additive loci determining the quantitative trait. (A) Simulations of a model without double reduction ( $\alpha_4 = 0$ ). (B) Simulations of a model with maximal double reduction ( $\alpha_4 = 1/6$ ) (for all loci in the finite L simulations). We show window-smoothed values for visual clarity, with observed genetic variances averaged in windows of 20 generations every 10 generations. The black line marks  $e^{-t/4N}$ . We assume N = 250 and  $V_z(0) = 1$ . Alt text: "The decay of genetic variance in autotetraploid populations for a quantitative trait determined by L loci."

We model assortative mating by ploidy level in a similar way, assuming that a fraction  $\rho_{c_i}$  of the ovules of individual *i* are fertilized by pollen sampled from the  $c_i$ -ploid portion of the population, while a fraction  $1 - \rho_{c_i}$  is fertilized by pollen randomly sampled from the entire population.

$$w_{i} = \rho_{c_{i}} \frac{1}{N_{c_{i}}} \sum_{j=1}^{N} \delta_{c_{i},c_{j}} w_{ij} + (1 - \rho_{c_{i}}) \left[ \frac{1}{N} \sum_{j=1}^{N} w_{ij} \right]$$
(14)

<sup>258</sup> Where  $\delta_{x,y}$  is the Kronecker delta function.

## 259 **Results**

The results section is organized as follows: first we verify the correctness and accuracy of 260 the infinitesimal model for mixed-ploidy populations by comparison against simulations of 261 a model with finitely many loci. Next we assess how autotetraploid genetics and different 262 assumptions on the genetic variance of neotetraploids affect the probability of tetraploid 263 establishment in a marginal habitat relative to diploid establishment, starting from a 264 single maladapted migrant individual. We then study how continuous migration from a 265 predominantly diploid maladapted source population affects the relative establishment 266 probability, assessing the impact of migration load and MCE. Finally, we consider the 267 impact of prezygotic isolation mechanisms on tetraploid establishment in the marginal 268 habitat. Throughout, we assume  $\beta_3 = \sqrt{2/3}, \beta_4 = \sqrt{1/2}$  and  $\alpha_2 = \alpha_3 = \alpha_4 = 0$ , unless 269 stated otherwise. Note that while the former assumption is a natural default, the latter 270 is not easy to motivate, and we do investigate the impact of the  $\alpha$  parameters in some 271 detail below. 272

#### <sup>273</sup> Autotetraploid and mixed-ploidy infinitesimal model

We evaluate the accuracy of the autotetraploid infinitesimal model as an approximation 274 to the evolution of a quantitative trait determined by L additive loci. We find that the 275 infinitesimal model with inbreeding generally yields accurate predictions for the evolution 276 of the genetic variance when the number of loci is sufficiently large ( $L \ge 100$ , say, figs. 1) 277 and S1). Furthermore, we confirm that, in the absence of double reduction, the decay 278 in genetic variance due to inbreeding after a time t is well-predicted by  $e^{-t/4N}$  (fig. 1A), 279 as expected from the results of Arnold et al. (2012). As predicted, double reduction 280 (i.e.  $\alpha_4 > 0$ ) leads to an immediate increase in genetic variance (as it increases the 281 segregation variance), but leads to accelerated inbreeding, causing faster loss of variation 282 in the long term (figs. 1 and S1). Simulations for the mixed-ploidy model further confirm 283 the correctness of our infinitesimal approximation (fig. S2). 284

It is worth noting that, although inbreeding is slower in autotetraploids than in 285 diploids for the same population size, the tetraploid fraction of a diploid-dominated 286 mixed-ploidy population will have an equal or higher average inbreeding coefficient (fig. S3) 287 This is because in such a population, triploid and tetraploid individuals mostly arise 288 from gametes formed by diploid individuals, or by polyploid individuals with very recent 289 diploid ancestry (on average 1 + u + 2v generations ago for tetraploids, and  $1 + \frac{2}{3}(u + 2v)$ 290 generations ago for triploids, see supplementary section S2.3). A nonzero probability of 291 producing IBD diploid gametes ( $\alpha_k > 0$ ) will then further increase the inbreeding coef-292 ficient in the tetraploid and triploid fraction of the population relative to their diploid 293 progenitors (fig. S3). Therefore, as long as diploids dominate, harboring some fraction of 294 the gene pool in polyploid individuals has a negligible effect on the rate of inbreeding in 295 the mixed-ploidy population as a whole, and we find that the evolution of the inbreeding 296 coefficient over time is well predicted by  $1 - e^{-t/2N_e}$ , where the inbreeding-effective pop-297 ulation size is, to first order in u, given by (1-2u)N (supplementary section S2.4). This 298 is just the expected number of diploid individuals (to first order in u), highlighting that 299 when diploids dominate, polyploids do not contribute to the effective population size. 300

#### <sup>301</sup> Establishment from a single individual

Having established the validity of the mixed-ploidy infinitesimal model, we now use it to study the establishment of polyploids in a marginal habitat to which migrants from a mixed-ploidy source population are maladapted.

We first consider the establishment of a population from a single migrant individual 305 with trait value  $z_0 = 0$ . We assume u = 0 (i.e. there are no unreduced gametes, and 306 hence no newly formed polyploids) and compare the probability of establishment when 307 the migrant is diploid vs. tetraploid (fig. 2). For a given mixed-ploidy model (character-308 ized by parameters  $\alpha, \beta, u$  and v), the establishment probability depends on  $\gamma, \theta$  and V 309 through two dimensionless parameters,  $\gamma\sqrt{2V}$  and  $\theta/\sqrt{2V}$  (Barton and Etheridge, 2018), 310 corresponding to the intensity of selection and the degree of maladaptation, respectively. 311 We shall scale our results accordingly, assuming 2V = 1 throughout. 312

For a fixed degree of maladaptation  $\theta$ , the probability of establishment depends in a complicated way on the strength of selection. To see this, note that the expected number of offspring of an initial migrant of ploidy level k is  $e^{-\gamma\theta+\gamma^2k\beta_k^2V/4}$ , and the expected trait value among its offspring will be  $\gamma k\beta_k^2 V/2$ . A higher intensity of selection ( $\gamma$ ) therefore yields a stronger effect of initial maladaptation, but also causes a stronger response in the mean trait value. If the genetic variance is not constant across cytotypes (i.e.  $\beta_4^2 \neq 1/2$ ),



Figure 2: (A) Probability of establishment (defined as reaching N = 100) from a single diploid or tetraploid individual with trait value z = 0 for increasing selection intensity  $\gamma$ , for two different values of  $\theta$  (degree of maladaptation). We assume m = 0 and u = 0, i.e. there is no migration, and no unreduced gametes are produced. The trait is scaled in tetraploids so as to yield the same genetic variance at HWLE ( $\beta_4^2 = 1/2$ ). Note that when  $\gamma = 0$ , we obtain a critical branching process with a Poisson offspring distribution, so that the probability to reach N = 100 is  $\sim 1/100$  (Barton and Etheridge, 2018). (B) Probability of a tetraploid individual with trait value z = 0 successfully founding a population ( $P_4$ ), relative to the probability for a diploid individual with the same trait value ( $P_2$ ). The vertical dashed line marks  $\beta_4^2 = 1/2$ . (C) Probability of tetraploid establishment relative to the probability of diploid establishment (on a  $\log_{10}$  scale) across a range of values for  $\gamma$  and  $\beta_4$  ( $\theta = 2.5$ ). The vertical line again marks  $\beta_4^2 = 1/2$ . All results are estimated from 1.000.000 (A&B) or 500.000 (C) replicate simulations. Alt text: "The probability that a tetraploid or diploid population is established in the marginal habitat when there is a single founder individual depends on detailed quantitative genetic assumptions."

this response will differ for different ploidy levels. Different rates of inbreeding due to differences in ploidy level will then further cause rates of adaptation to differ, leading to different establishment probabilities, even when  $\beta_4^2 = 1/2$ .

Indeed, we find that reduced inbreeding in tetraploids substantially increases the like-322 lihood of tetraploid establishment relative to diploids across a large part of the parameter 323 space (fig. 2A). For the  $\beta_4^2 = 1/2$  case, the establishment probability for tetraploids can 324 be more than ten times as high as for diploids depending on the selection gradient  $(\gamma)$  and 325 the degree of maladaptation ( $\theta$ ) (fig. 2A). As the segregation variance and initial trait 326 value are the same across these simulations, this is a consequence only of the reduced 327 rate of inbreeding, which slows down the exhaustion of the genetic variance carried by 328 the initial migrant individual. Evidently, the scaling of the genetic variance across ploidy 329 levels has a profound effect on the establishment probability, but only when  $\beta_4$  is close 330 to 0.5 (i.e. individual alleles have almost half the effect size in tetraploids compared to 331 diploids) is the benefit of the slower rate of inbreeding in tetraploids canceled by the 332 reduction in the equilibrium genetic variance (fig. 2B,C). 333

#### <sup>334</sup> Establishment with recurrent migration

We next consider establishment in the new habitat when there is a continuous influx of migrants (*m* migrants per generation on average) coming from a large, noninbred and predominantly diploid source population at cytotype equilibrium. In this setting, establishment is certain to happen eventually, and we are interested in the probability that a tetraploid population establishes before a diploid one does.

We hypothesized that two counteracting processes affect the probability of autotetraploid establishment in this scenario. On the one hand, increased migration will increase the probability that an otherwise likely successful tetraploid migrant suffers from MCE in the early generations while the population size is low, because migrants are likely



Figure 3: Establishment with recurrent migration. (A) Expected time until a population is established in the marginal habitat for increasing rates of migration and different degrees of maladaptation ( $\theta$ ). Results are shown for the case with  $\alpha_k = 0$  for k = 2, 3, 4. (B) Proportion of simulation replicates in which tetraploids established. The dots connected by solid lines show simulation results with  $\alpha_k = 0$ , whereas the dots connected by dashed lines show simulation results with  $\alpha_2 = 1/2, \alpha_3 = 1/4$  and  $\alpha_4 = 1/6$  (i.e. maximum  $\alpha$ ). The horizontal lines mark the establishment probabilities in the limit as  $m \to 0$  (solid lines: without double reduction; dashed lines: maximum  $\alpha$ ). The black horizontal line marks the proportion of tetraploid migrants (i.e. the proportion of tetraploids at equilibrium in the source population,  $\approx 0.3\%$ ). The baseline predictions (horizontal lines) are based on 500.000 simulation replicates. All other results are based on 100.000 replicate simulations. We assume  $\gamma = 0.25$  and u = v = 0.05. Alt text: "The likelihood of tetraploid establishment depends on the migration rate when there is recurrent migration."

to be diploid. On the other hand, tetraploids are more strongly reproductively isolated 344 from a typical migrant, so that a tetraploid subpopulation should be less prone to mal-345 adaptive gene flow. Hence, conditional on evading MCE, they should be able to adapt 346 to the new habitat at a rate which is not strongly affected by the migration rate. This 347 contrasts with diploids, which interbreed freely with maladapted migrants, resulting in a 348 pulling back of the trait mean towards that of the source population. Lastly, as the mean 349 trait value on the island increases in diploids during adaptation, tetraploid offspring will 350 have more extreme phenotypes on average than diploid offspring when  $\beta_4 > 1/2$ , which 351 may also aid their establishment (irrespective of m). 352

As expected, we find that the time to establishment (of a population of either ploidy 353 level) first decreases with increasing migration as a result of a larger influx of potentially 354 succesful migrants, but later increases with increasing migration due to swamping by 355 gene flow (fig. 3A). Importantly, the tetraploid establishment probability is considerably 356 larger than the expected proportion of tetraploid migrants over almost the entire param-357 eter range examined (fig. 3B, black dashed line). However, the probability of tetraploid 358 establishment does decline monotonically with the migration rate, showing that the neg-359 ative effects of MCE on tetraploid establishment outweigh the positive effects of reduced 360 maladaptive gene flow in the absence of prezygotic isolation. 361

Our simulations further show that the mechanism of unreduced gamete formation (as determined by the  $\alpha_2$  parameter) can affect the establishment probability (fig. 3B, dashed lines). This is mainly because the phenotypic variance of a newly formed tetraploid is increased by a factor  $(1 + \alpha_2)$ , thereby increasing the chance that a tetraploid migrant is well-adapted to the marginal habitat. The rate of double reduction ( $\alpha_4$ ) has a more limited effect (fig. S4).

Established diploid populations are more inbred on average than established tetraploids when migration is weak, but the difference is slight except when there is no migration at



Figure 4: Evolution of the mean inbreeding coefficient and trait value across simulation replicates where diploids (blue) or tetraploids (orange) established eventually. Average F and z by population size are shown for increasing rates of migration (m) from the predominantly diploid source population. All results are based on 1000 succesful establishment replicates. We assume equal equilibrium variance across ploidy levels and  $\gamma = 0.25$ ,  $\theta = 2$ , 2V = 1 and u = v = 0.05. For the m = 0 simulations, the trait value of the initial migrant was Gaussian with mean zero and variance 2V, and u = v = 0 is assumed. Alt text: "How the population evolves during the establishment process depends on the rate of migration and can be quite different for the different cytotypes."

all (fig. 4, top row). For stronger migration (m > 0.1), the opposite holds. This is a result of two interacting processes. On the one hand, inbreeding is slower in tetraploids, so that during adaptation and establishment from a single or limited number of outbred individuals, the inbreeding coefficient is expected to increase less rapidly. On the other hand, migration mostly introduces unrelated diploids, which cross more readily with diploids than tetraploids, reducing the average relatedness more strongly in established diploid than in tetraploid populations.

Conditional on establishment, tetraploids have a higher trait mean than diploids 377 (fig. 4, bottom row). In the absence of migration, this is a consequence of the reduced 378 rate of inbreeding and the resulting increased adaptive potential of tetraploids. For weak 379 migration, the difference in trait values between diploids and tetraploids, conditional on 380 eventual establishment, is limited. This indicates the beneficial effects of migration on 381 establishment in diploids: migration introduces new variation on which selection can act, 382 counteracting the loss of genetic variance due to inbreeding. The genetic variance con-383 tributed by migration is however negligible in tetraploids. When migration is strong, 384 tetraploids have markedly larger trait values than diploids (m = 1 in fig. 4), showing that 385 diploids suffer strongly from maladaptive gene flow when the population size is low, while 386 tetraploids are much more reproductively isolated from migrants. Furthermore, in these 387 replicates, tetraploids tend to emerge and rise in frequency at larger population sizes on 388 the island, and hence tend to derive from diploids that already experienced several gen-389 erations of selection. These neotetraploids, deriving from diploid parents with z > 0, will 390 have more extreme phenotypes on average (see methods) and hence be better adapted. 391

#### <sup>392</sup> Loss of self-incompatibility, selfing and assortative mating

<sup>393</sup> When polyploidization disrupts an existing SI system (see e.g. Robertson et al. (2011); <sup>394</sup> Zenil-Ferguson et al. (2019); Novikova et al. (2023)), we expect that tetraploids suffer



**Figure 5:** (A) Establishment with recurrent migration and selfing in polyploids. The solid lines show the case where diploids are self-incompatible The dashed transparent lines show the case where diploids do random self-fertilization (i.e. self-fertilization occurs with probability 1/N), Triploids and tetraploids have the same selfing rate.  $\sigma = 0.0$  refers to random self-fertilization. (B) Establishment with recurrent migration and assortative mating by cytotype. The rate of assortative mating is determined by  $\rho_k = \rho$ for k = 2, 3, 4, where  $\rho_k$  is the probability that an ovule from a k-ploid mother is pollinated by a kploid father. All results are based on 50.000 replicate simulations. We assume  $\gamma = 0.25, \theta = 2$  and u = v = 0.05. Alt text: "Prezygotic isolation and reproductive assurance due to selfing and assortative mating have a strong effect on the probability that a tetraploid population establishes in the marginal habitat."

less from MCE, as some portion of their ovules are now assured to be fertilized by diploid 395 gametes, irrespective of the composition of the population. At the same time, we expect 396 that accelerated inbreeding in selfing tetraploids diminishes the adaptive advantage of 397 tetraploids. We find that when polyploidization is associated with the loss of a SI system 398 (i.e. when diploids are self-incompatible, but tetraploids are not), tetraploids have a 399 strongly increased establishment probability (fig. 5). This is the case even when the selfing 400 rate  $\sigma$  in tetraploids is zero (in which case, under our modeling assumptions, there is only 401 random selfing, i.e. the *realized* selfing rate in tetraploids is 1/N). Furthermore, we find 402 that when the selfing rate is sufficiently high ( $\geq 0.4$  in fig. 5A), the relative establishment 403 probability of tetraploids increases with increasing migration rate. In this regime, the 404 effects of migration on MCE and reproductive assurance in diploids are compensated by 405 the stronger maladaptive gene flow experienced by diploids. 406

Self-incompatibility is clearly a strong disadvantage when colonizing a novel habitat, 407 as a self-incompatible population of size one can never reproduce. However, even when 408 diploids are self-compatible, polyploids may still have increased rates of self-fertilization 409 (for instance due to altered flower morphology). For the sake of comparison, fig. 5A 410 also shows results where diploids are assumed to be self-compatible with  $\sigma_2 = 0$  (i.e. 411 random selfing, dashed transparent lines). The tetraploid establishment probability is still 412 markedly increased when  $\sigma_4 \geq 0.4$ , and as for the simulations with self-incompatibility, 413 migration still promotes the probability of tetraploid establishment when the selfing rate 414 in polyploids is sufficiently large compared to the diploid selfing rate. 415

Another prezygotic isolating mechanism that has often been considered relevant for explaining tetraploid establishment is assortative mating by ploidy level, where ovules from a tetraploid are more likely to be fertilized by pollen coming from a tetraploid – irrespective of the trait values of these individuals. Clearly, assortative mating increases the probability of tetraploid establishment (fig. 5B), although not as strongly as the loss of an SI system does. Again, we find that for some parameter values (roughly  $\rho \geq 0.4$ ), assortative mating may be strong enough so that tetraploid establishment increases with increasing migration rates, suggesting that tetraploids evade maladapive gene flow sufficiently to overcome MCE. Note that the case  $\rho = 1$  amounts to complete production isolation

<sup>425</sup> prezygotic isolation.

## 426 Discussion

The observation that polyploid populations tend to inhabit more extreme habitats or 427 occur at the edge of the range of their conspecific diploids has spurred considerable 428 interest among botanists and evolutionary biologists (Kolář et al., 2017; Rice et al., 2019; 429 Van de Peer et al., 2021; Griswold, 2021; Mortier et al., 2024). An important question is 430 whether such patterns emerge because polyploids are somehow more tolerant to extreme 431 environmental conditions (i.e. they somehow are intrinsically more fit than diploids in 432 marginal habitats), or whether other aspects of the population dynamics of mixed-ploidy 433 populations may favor the establishment of polyploid subpopulations. 434

In this study, we worked out the infinitesimal model for an additive polygenic trait in autotetraploids and mixed-ploidy populations and used it to study the establishment of tetraploids in a marginal habitat by means of individual-based simulations. Assuming the trait to be under directional selection in the marginal habitat, and migration of maladapted individuals from a predominantly diploid source, we sought to determine under which conditions tetraploids are more likely to establish a stable population.

Throughout, we have assumed a relatively high and constant rate of unreduced gamete formation u and triploid fertility v in all our simulations (5%), whereas these are known to be variable across the population, and at least in part genetically determined (Kreiner et al., 2017a; Clo et al., 2022). We ignore such complications, and hence do not take the actual establishment probabilities very serious, focusing instead on how migration load and prezygotic isolation affect the tetraploid establishment probability.

Similarly, we have ignored mutation, which would reduce the rate at which genetic 447 variation is lost through inbreeding (Barton et al., 2017), and would likely do so differently 448 across cytotypes (i.e.  $\mu V_m$  is expected to differ for different ploidy levels, where  $\mu$  is the 449 mutation rate and  $V_m$  the mutational variance). The contribution of new mutation to the 450 genetic variance on the timescales we consider should however be very limited. Indeed, 451 any individual at the time of establishment derives from a completely outbred migrant 452 individual a relatively short time in the past, so that the opportunity for mutation to 453 contribute to differences in establishment probability between diploids and tetraploids is 454 negligible for realistic  $\mu V_m$ . 455

Importantly, we assumed no intrinsic advantage or disadvantage of polyploids in the 456 marginal habitat, i.e. the expected fitness of a migrant individual is the same regardless 457 of the ploidy level. Differences in the likelihood of polyploid establishment are hence 458 caused solely by aspects of autopolyploid genetics and the barrier to gene flow between 459 subpopulations of different ploidy levels. This is undoubtedly unrealistic, as both trait 460 values and fitness will often differ systematically across ploidy levels (see e.g. Porturas 461 et al. (2019)). For instance, neopolyploids are likely to suffer intrinsic fertility issues 462 due to meiotic irregularities associated with multivalent formation (Bomblies et al., 2016; 463 Novikova et al., 2023), and triploids may be inviable due to issues with endosperm devel-464 opment (Bretagnolle and Thompson, 1995). 465

Similarly implausible is the assumption of a constant equilibrium genetic variance across cytotypes ( $\beta_4^2 = 1/2$  in our model), which we used in most of our results (but see

fig. 2). Empirical data on how the genetic variance scales across ploidy levels is scant and 468 suggests that there is no general rule (Gallais, 2003; Porturas et al., 2019). The meta-469 analysis performed by Porturas et al. (2019) does indicate that trait variance across ploidy 470 levels is often fairly constant, so the assumption of equal genetic variance is arguably a 471 reasonable default. It should be noted however that other authors have made different 472 assumptions on how allelic effects (and hence genetic variance) scale across ploidy levels 473 (in particular Griswold (2021), who scaled allelic effects in a way that is equivalent to 474  $\beta_4 = 1/2$  in our model). Such assumptions evidently impact the likelihood of polyploid 475 establishment (fig. 2). More empirical data on quantitative traits in experimental or nat-476 ural mixed-ploidy populations is needed to assess whether the mixed-ploidy infinitesimal 477 model can adequately describe the genetics of quantitative traits across cytotypes, and 478 to suggest plausible values for the relevant parameters  $(\alpha, \beta)$ . 479

When migration is weak, successful establishment is not affected by maladaptive gene 480 flow and we can treat establishment in the marginal habitat as independent trials of 481 founding a population from a single individual. In order to avoid extinction, the popu-482 lation has to increase the trait mean by a sufficient amount before the genetic variation 483 carried by the initial migrant individual is exhausted. The probability that the population 484 manages to do so depends on the degree of maladaptation, the intensity of selection and 485 the rate of inbreeding. We find that the decreased rate of inbreeding in autotetraploids 486 gives a rare tetraploid migrant a larger adaptive potential than a diploid migrant, even 487 if the genetic variance carried by the founding individual is the same. 488

In the presence of maladaptive gene flow, a nascent tetraploid subpopulation suffers 489 from MCE, and although polyploids are more reproductively isolated from a typical 490 migrant (and hence suffer less maladaptive gene flow), MCE will increasingly hamper the 491 establishment of tetraploids as the rate of migration grows. Nevertheless, it is important 492 to remark that despite MCE, the probability of tetraploid establishment in the marginal 493 habitat can be an order of magnitude higher than expected based on the frequency of 494 tetraploid migrants (i.e. is roughly of order u instead of  $u^2$ ) when migration is sufficiently 495 weak and maladaptation sufficiently high. 496

Additional sources of prezygotic isolation such as selfing and assortative mating by 497 cytotype may further boost the probability of tetraploid establishment. These processes 498 interact with the rate of migration, so that when selfing/assortative mating occurs above 499 some threshold rate, the tetraploid establishment probability increases with increasing 500 migration rates, whereas below the threshold it decreases with increasing migration pres-501 sure. In the latter case, the advantage that tetraploids have when it comes to avoiding 502 maladaptive gene flow is not strong enough to overcome the effects of MCE, whereas in 503 the former case it is. 504

A major weakness of the present work, and an important caveat, is that we have 505 ignored inbreeding depression and dominance throughout. Including dominance in the 506 infinitesimal framework is already challenging for diploids (requiring the tracking of four-507 way identity coefficients; Barton et al. (2023)), and appears intractable for higher ploidy 508 levels. However, autopolyploidy has important consequences whenever dominance is rel-509 evant, as in the case of inbreeding depression (Ronfort, 1999; Gallais, 2003; Husband 510 et al., 2008; Clo and Kolář, 2022). Indeed, when inbreeding depression is due to recessive 511 deleterious variation, it is expected to be less expressed in neotetraploids because homozy-512 gous genotypes should be much rarer than in their diploid parents (the 'masking' effect; 513 Husband and Schemske (1997); Otto and Whitton (2000)). Inbreeding during the estab-514 lishment process should therefore incur a higher fitness cost in diploids relative to neote-515

traploids, and hence further increase the probability of tetraploid establishment. How this plays out depends however on the *rate* at which populations become inbred, which will differ between cytotypes and will depend strongly on the mating system. In outcrossing populations, inbreeding occurs at a slower rate in tetraploids, further decreasing inbreeding depression and aiding tetraploid establishment. However, when polyploidization is associated with increased selfing (as when it disrupts an existing SI system), increased inbreeding depression in autotetraploids may prevent their establishment.

Dominance and inbreeding depression may strongly affect the complicated interaction 523 between selfing and migration load in determining tetraploid establishment. Griswold 524 (2021) studied the case where local fitness is determined by a single biallelic locus, and 525 investigated the interaction between inbreeding depression and migration load (where 526 inbreeding depression is modeled as a fixed fitness reduction in offspring produced by 527 selfing). In his model, inbreeding depression is different between cytotypes (assuming 528 stronger inbreeding depression in diploids), so that tetraploids are able to produce more 529 offspring through selfing relative to diploids, who have to rely more on outcrossing. How-530 ever, outcrossing incurs maladaptive gene flow, and thereby puts the diploids at a disad-531 vantage. He found that autotetraploids can establish when adaptation in the peripheral 532 habitat is conferred by recessive alleles (so that migration load is expressed when migrant 533 alleles are rare) and when inbreeding depression in tetraploids is lower than in diploids. 534 It would be very interesting to combine the infinitesimal framework with some form of 535 inbreeding depression to investigate in a more realistic model whether the combination of 536 maladaptive migration and differential inbreeding depression could explain the prevalence 537 of polyploid subpopulations at range edges. 538

In the long term, polyploids are expected to accumulate a larger mutation load when 539 deleterious variation is recessive due to less efficient purging, and this may yield *increased* 540 inbreeding depression (Vlček et al., 2025). These effects have been studied in the context 541 of range expansions (Booker and Schrider, 2024). However, this applies only to poly-542 ploids that have been established for a long time. In our case, polyploids are always 543 recently descended from diploid ancestors, and they will not have accumulated more 544 deleterious mutations than their diploid counterparts, so that polyploidy should lead to 545 reduced rather than increased inbreeding depression when selfing rates are similar (as dis-546 cussed above). Interestingly, the interplay between the effects of polyploidy on different 547 timescales could yield an equilibrium situation that may characterize many mixed-ploidy 548 populations in nature: although sometimes polyploids could enjoy enhanced establish-549 ment probabilities in peripheral habitats, the accumulation of mutational load may in the 550 long-term limit further range expansion or even lead to competitive exclusion by diploids. 551 Further modeling efforts could provide more insights into the plausibility of such a model. 552

While in this study we focused on polyploid establishment in a peripheral habitat and 553 how this is affected by migration from a diploid source, the mixed-ploidy infinitesimal 554 framework could be used to address many other eco-evolutionary questions that arise 555 in the study of mixed-ploidy populations. For instance, it could be of interest to de-556 velop an individual-based model along the lines of Oswald and Nuismer (2011) to study 557 competitive exclusion versus coexistence of cytotypes in mixed-ploidy populations with 558 partial selfing, assortative mating, competition, etc., but where the focal trait that de-559 termines the various fitness components is not controlled by a few large-effect loci (as in 560 Oswald and Nuismer (2011)), but many loci of small effect. Similarly, our model could 561 be straightforwardly extended to include population regulation and stabilizing selection, 562 which would allow us to study polyploid establishment along an environmental gradi-563

ent and the potential of polyploidization to promote range expansions (Polechová and Barton, 2015)

<sup>565</sup> Barton, 2015).

# <sup>566</sup> Supplementary material

567 Supplementary material is available online at *Evolution*.

## 568 Data availability

- <sup>569</sup> Individual-based simulations for the mixed-ploidy infinitesimal model were implemented
- in Julia (Bezanson et al., 2017). Documented code and simulation notebooks are available at https://github.com/arzwa/InfGenetics (10.5281/zenodo.15560879).
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## 575 Conflict of interest

576 I have no conflict of interest to declare.

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